

A New Method for Analyzing Time Intensity Curves

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Outline

- A novel Modelling approach is introduced and parameters are estimated via an EM algorithm. Smoothing splines are also aggregated.
- Four simulations are performed on simulated data; we obtain fitted curves based on the assumptions of homoscedastic and heteroscedastic error terms, respectively, at each time point.
- Real fruit liqueur data are analyzed.
- Discussion and suggestions for future work.

Aim

- To estimate underlying time intensity curves and cluster individuals.
- How it can help us to discover useful information about attributes.

Modelling Framework

- TI curves are monotonically increasing until time T_{\max} and then monotonically decreasing thereafter.

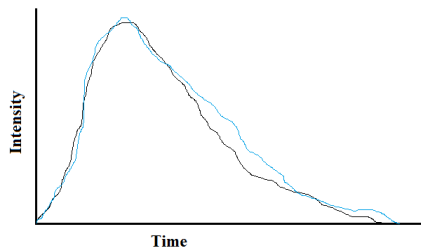


Figure 1: TI Curves

- We represent this dependence with a Markovian error term.

Modelling Framework

- Let z_i be the observed TI value and x_i be the latent TI value so that

$$z_i = \begin{cases} \max\{x_{i-1}, x_i\} & \text{for } i = 2, \dots, k, \\ \min\{x_{i-1}, x_i\} & \text{for } i = k + 1, k + 2, \dots, n. \end{cases}$$

where $X_i \sim \mathcal{N}(\mu_i, \sigma_i^2)$, k is T_{\max} and n is total time points.

For the Markovian error term, we consider two options:

- 1 Homoscedasticity: there is a common standard deviation across all time points for all panellists, i.e., $\sigma_i^2 = \sigma^2$, for $i = 1, 2, \dots, n$.
- 2 Heteroscedasticity: each time point has its own standard deviation.

Modelling Framework

- The complete-data log-likelihood function using the homoscedastic σ is

$$\mathcal{L}(\mu_1, \dots, \mu_n, \sigma \mid z_1, \dots, z_n) = -\frac{np}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n \sum_{j=1}^p ((X_{ij} - \mu_i)^2 \mid \mathbf{Z}),$$

- The complete-data log-likelihood function using the heteroscedastic σ_i is

$$\mathcal{L}(\mu_1, \dots, \mu_n, \sigma_1, \dots, \sigma_n \mid z_1, \dots, z_n) = -\frac{np}{2} \log(2\pi\sigma_i^2) - \frac{1}{2\sigma_i^2} \sum_{i=1}^n \sum_{j=1}^p ((X_{ij} - \mu_i)^2 \mid \mathbf{Z}).$$

EM Algorithm

- The EM algorithm is an iterative method for finding maximum likelihood estimates of parameters where there are unobserved or missing data.
- An expectation (E-) step that computes the expectation of the complete-data log-likelihood given the current estimates is followed by a maximization (M-) step wherein the expectation of the complete-data log-likelihood is maximized with respect to the model parameters.
- The E- and M-steps are iterated until convergence.

Truncated Normal Distribution

- $\mathbf{X} \mid \mathbf{Z} \sim$ truncated $\mathcal{N}(\mu_i, \sigma^2)$
- $\mathbf{X} \mid \mathbf{Z} \sim$ truncated $\mathcal{N}(\mu_i, \sigma_i^2)$

so, the expectation

$$\mathbb{E}(X \mid a < Z < b) = \mu + \frac{\phi(\frac{a-\mu}{\sigma}) - \phi(\frac{b-\mu}{\sigma})}{\Phi(\frac{b-\mu}{\sigma}) - \Phi(\frac{a-\mu}{\sigma})} \sigma,$$

and the variance

$$\text{Var}(X \mid a < Z < b) = \sigma^2 \left[1 + \frac{\frac{a-\mu}{\sigma} \phi(\frac{a-\mu}{\sigma}) - \frac{b-\mu}{\sigma} \phi(\frac{b-\mu}{\sigma})}{\Phi(\frac{b-\mu}{\sigma}) - \Phi(\frac{a-\mu}{\sigma})} - \left(\frac{\phi(\frac{a-\mu}{\sigma}) - \phi(\frac{b-\mu}{\sigma})}{\Phi(\frac{b-\mu}{\sigma}) - \Phi(\frac{a-\mu}{\sigma})} \right)^2 \right]$$

E step

$$\mathbb{E}(X_i | z_{i-1}, z_i) = \begin{cases} z_i & \text{if } z_i > z_{i-1}, \\ \mu_i - \frac{\phi(\frac{z_i - \mu_i}{\sigma})}{\Phi(\frac{z_i - \mu_i}{\sigma})} \sigma & \text{if } z_i = z_{i-1}, \end{cases}$$

for $i = 2, \dots, k$ and

$$\mathbb{E}(X_i | z_{i-1}, z_i) = \begin{cases} z_i & \text{if } z_{i-1} > z_i, \\ \mu_i + \frac{\phi(\frac{z_i - \mu_i}{\sigma})}{\Phi(\frac{z_i - \mu_i}{\sigma})} \sigma & \text{if } z_{i-1} = z_i, \end{cases}$$

for $i = k + 1, \dots, n$.

M step

Under the homoscedastic assumption, the expected value of the complete-data log-likelihood is given by

$$Q_1(\mathbf{X}, \mathbf{Z} \mid \boldsymbol{\mu}, \sigma^2) = -\frac{np}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n \sum_{j=1}^p E\{(X_{ij} - \mu_i)^2 \mid \mathbf{Z}\},$$

where p is number of repetitions, n is number of time points and $\boldsymbol{\mu}$ is a $n \times 1$ matrix.

$$\hat{\mu}_i = \frac{1}{p} \sum_{j=1}^p E\{X_{ij} \mid \mathbf{Z}\} \quad \text{and} \quad \hat{\sigma}^2 = \frac{1}{np} \sum_{i=1}^n \sum_{j=1}^p E\{(X_{ij} - \mu_i)^2 \mid \mathbf{Z}\}.$$

M step

For the second assumption-heteroscedastic σ , the expected value of the complete-data log-likelihood function is

$$Q_2(\mathbf{X}, \mathbf{Z} \mid \boldsymbol{\mu}, \sigma^2) = -\frac{p}{2} \sum_{i=1}^n \log \sigma_i^2 - \sum_{i=1}^n \frac{1}{2\sigma_i^2} \sum_{j=1}^p E \{ (X_{ij} - \mu_i)^2 \mid \mathbf{Z} \} + C,$$

where C is a constant.

$$\hat{\mu}_i = \frac{1}{p} \sum_{j=1}^p E \{ X_{ij} \mid \mathbf{Z} \} \quad \text{and} \quad \hat{\sigma}_i^2 = \frac{1}{p} \sum_{j=1}^p E \{ (X_{ij} - \mu_i)^2 \mid \mathbf{Z} \}.$$

Smoothing Spline: why

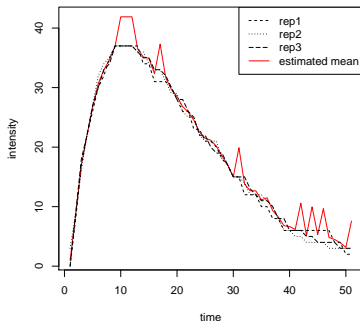


Figure 2: Fitted Curve

Smoothing Spline

- The penalized spline smoothing was introduced by O'Sullivan (1986).
- This smoothing method with flexible choice of bases and penalties can be viewed as a compromise between regression and smoothing splines which are piecewise polynomials with pieces smoothly connected together.

Smoothing Spline

- Let (x_i, Y_i) , so that $x_1 < x_2 < \dots < x_n$, be a sequence of observations modelled by the relation $Y_i = \mu(x_i)$. The penalized sum of squares is

$$S(\mu) = \sum_{i=1}^n (Y_i - \mu(x_i))^2 + \lambda \int_a^b \mu''(x)^2 dx,$$

- μ is any twice-differentiable function on $[a, b]$ and λ is a smoothing parameter.
- The first term measures the closeness of the fitted function to the data, while the second penalizes the curvature in the function.
- The smoothing spline estimate $\hat{\mu}$ of the function μ is

$$\hat{\mu} = \arg \min_{\mu \in \mu} S(\mu).$$

Preparation

- 1 Randomly generate the latent TI x_i values which follow a normal distribution with parameters μ_i and $\sigma = 0.01$.
- 2 A straightforward method to generate observed data z_1, \dots, z_n is given below:

$$z_1 = x_1, z_2 = \max(x_1, x_2), \dots, z_{k-1} = \max(x_{k-2}, x_{k-1}), z_k = \max(x_{k-1}, x_k), \\ z_{k+1} = \min(x_k, x_{k+1}), \dots, z_{n-1} = \min(x_{n-2}, x_{n-1}), z_n = \min(x_{n-1}, x_n),$$

where $n = 51$.

Preparation

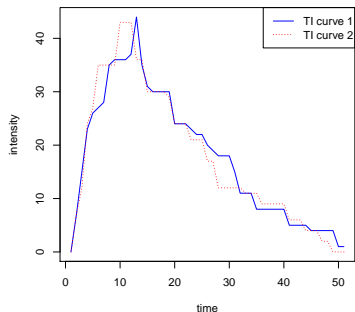


Figure 3: generate TI curves

Simulation Results: Homoscedastic Model

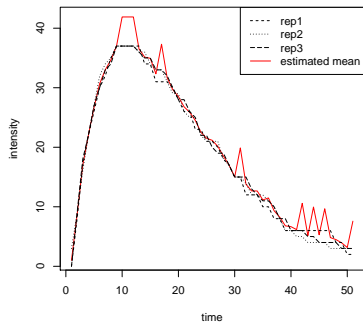
When $\sigma = 0.01$ 

Figure 4: Fitted Curve

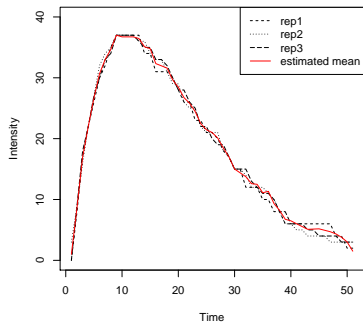


Figure 5: Smooth Curve

Simulation Results: Homoscedastic Model

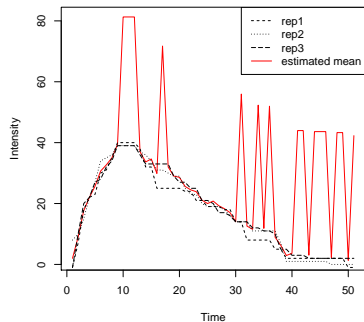
When $\sigma = 0.03$ 

Figure 6: Fitted Curve

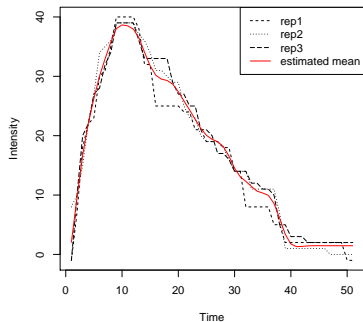


Figure 7: Smooth Curve

Simulation Results: Heteroscedastic Model

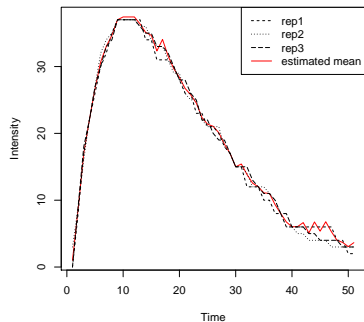
When $\sigma = 0.01$ 

Figure 8: Fitted Curve

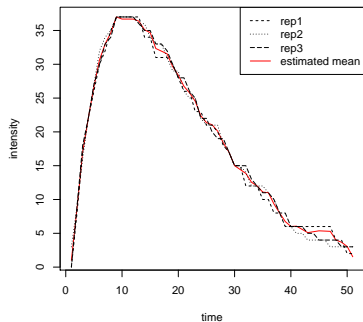


Figure 9: Smooth Curve

Simulation Results: Heteroscedastic Model

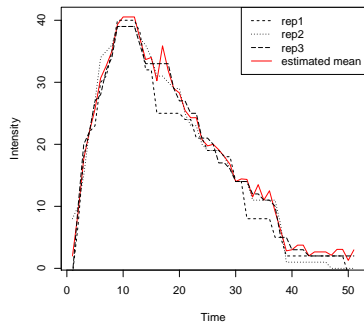
When $\sigma = 0.03$ 

Figure 10: Fitted Curve

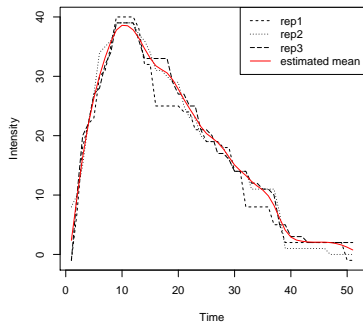


Figure 11: Smooth Curve

Results: Homoscedastic Model

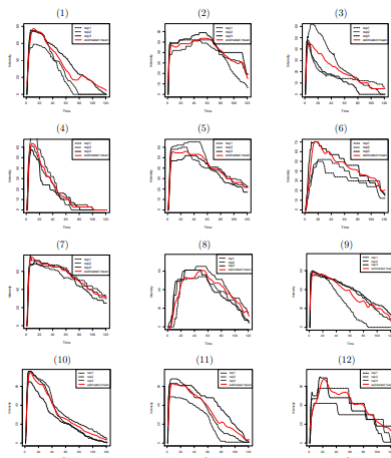


Figure 12: Smooth curves for product A

Results:Heteroscedastic Model

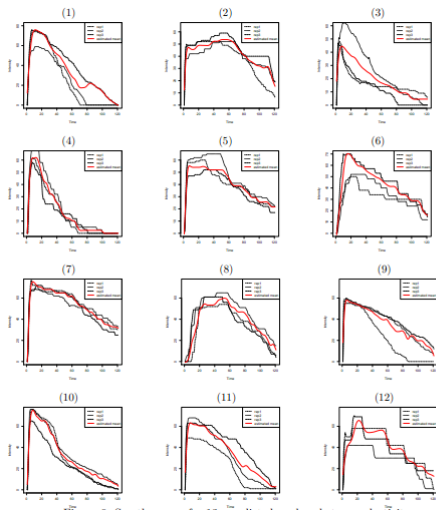


Figure 13: Smooth curves for product A

Clustering

- Group 1: panelist 1, 3, 10
- Group 2: panelist 2, 5, 6, 7, 12
- Group 3: panelist 9, 11
- Group 4: panelist 4
- Group 5: panelist 8

Results: Homoscedastic Model

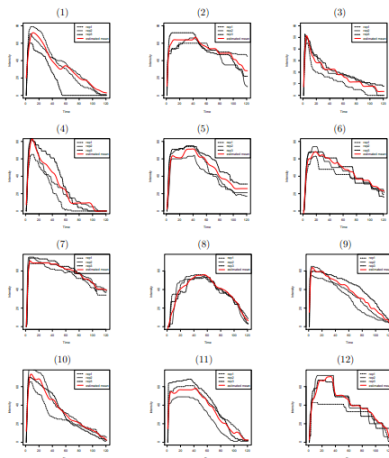


Figure 14: Smooth curves for product B

Product A vs. Product B: Homoscedastic Model

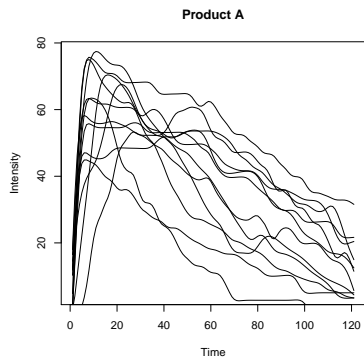


Figure 15: Smooth Curves for 12 panelists

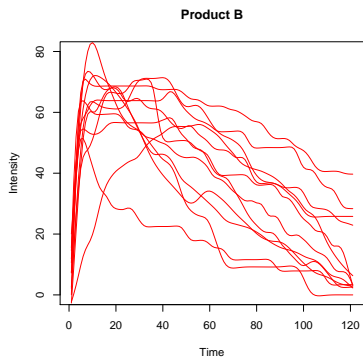


Figure 16: Smooth Curves for 12 panelists

Product A vs. Product B: Homoscedastic Model

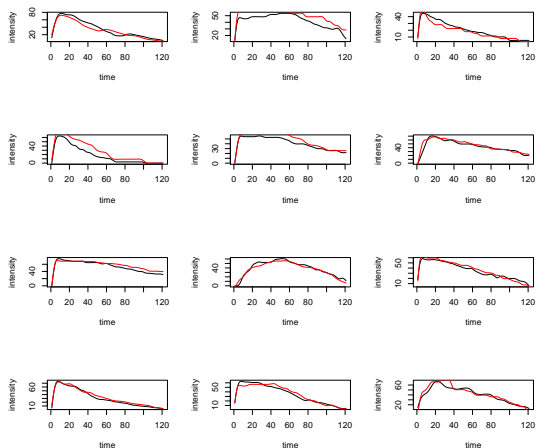


Figure 17: Smooth Curves between product A and B for each panelist

Conclusion

- Using different assumptions, the smoothing curves have similar shapes and are a representation of 3 TI curves.
- Recommending using homoscedastic σ obtain smooth TI curves.
- There is variation among the panelists for product A and product B.
- Panelists give very similar smoothing curves between product A and B.

Future Work

In the future, the problem of dealing with T_{max} . Because it is the crucial part of conducting a accurately fitted curve.

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